Common breast cancer susceptibility alleles and the risk of breast cancer for BRCA1 and BRCA2 mutation carriers: implications for risk prediction

Supplementary material

# Appendix 1

# **Genotyping Quality Control**

Genotyping was performed using either the iPLEX or Taqman platforms. To ensure genotyping consistency, included at least 2% of the samples in duplicate, no template controls in every plate, and a random mixture of affected and unaffected carriers. Samples that failed for two or more of the SNPs genotyped (or ≥20% of the SNPs typed if more than three SNPs were analysed in that genotyping round) were excluded from the analysis. The genotype data for a given SNP and a given study were included in the analysis only if the call rate was >95% after samples that failed at multiple SNPs had been excluded. The concordance between duplicates had to be at least 98%. To assess the accuracy of genotyping across genotyping centers, all centers genotyped 95 DNA samples from a standard test plate (Coriell Institute) for all SNPs. If the genotyping was inconsistent for more than one sample in the test plate, the study was excluded from the analysis of that SNP. As an additional genotyping quality-control check, we also evaluated the deviation from Hardy-Weinberg equilibrium (HWE) for unrelated subjects separately for each SNP and study. Studies with a HWE p-value of less than 0.005 were excluded from the analysis. If HWE p-values were in the range 0.005-0.05 we examined the genotyping cluster plots; none revealed any unusual patterns and these studies were therefore included in all the analyses. For the three novel polymorphisms investigated, one study failed the quality control criteria and was excluded from the analysis of the rs4973768 (SLC4A7/NEK10) polymorphism and five studies were excluded from the

analysis of the rs10941679 (5p12) SNP. No studies were excluded from the analysis of the rs6504950 (*STXBP4/COX11*) SNP.

A number of samples (3344 - 5150 mutation carriers depending on SNP) were regenotyped by iPLEX for the six previously published SNPs in a multiplex along with the new SNPs, as part of the effort to genotype additional samples from mutation carriers who were recruited subsequent to the original reports. Samples with inconsistent genotypes between the two genotyping rounds were excluded from the analysis of each polymorphism (discordance rate: 0.002-0.006 depending on the SNP).

# Appendix 2

To evaluate the combined effects of the significant SNPs on breast cancer risk, for each pair of SNPs we fitted models of the form:  $\lambda(t) = \lambda_0(t) \exp(\beta_1 x_1 + \beta_2 x_2 + \gamma x_1 x_2)$  where  $\beta_1$  is the per-allele log-hazard ratio for SNP1,  $\beta_2$  is the per-allele log-hazard ratio for SNP2,  $\gamma$  is a term for the interaction between SNP1 and SNP2 and  $x_1$ ,  $x_2$  represent the number of minor alleles at locus 1 and 2 respectively (e.g. 0,1,2). To test for interactions between a pair of SNPs, we tested for the significance of the interaction parameter (i.e.  $\gamma \neq 0$ ).

# Implications for risk predictions

We estimated the absolute risk of developing breast cancer based on the joint distribution of all SNPs that were significantly associated with risk for either *BRCA1* or *BRCA2* mutation carriers. To do this, we assumed that the average, age-specific breast cancer incidences for *BRCA1* and *BRCA2* mutation carriers, over all modifying loci, agreed with published penetrance estimates for *BRCA1* and *BRCA2* (1). Under this model, and assuming independence between the SNPs, the breast cancer incidence  $\lambda(t,X)$  at age t for a mutation carrier with genotype vector  $\mathbf{X}=(x_1, x_2, ..., x_n)$  at loci 1 to n, depends on the underlying genotypes according to :  $\lambda(t,X) = \lambda_0(t) \exp(\beta^t X)$  where  $\lambda_0(t)$  is the baseline incidence and  $\boldsymbol{\beta} = (\beta_1, \beta_2, ..., \beta_n)$  is the vector of logHRs estimates for SNPs 1 to n.  $\lambda_0(t)$  is unknown.  $\exp(\beta^t X)$  represents the combined HR across all loci. If  $\mu(t)$  is the average incidence at age t obtained from published studies, then:

$$\mu(t) = \frac{\sum_{i} p(X_i) \lambda(t, X_i) S(t-1, X_i)}{\sum_{i} p(X_i) S(t-1, X_i)}$$
 (1).

 $p(X_i)$  is the probability of multilocus genotype i  $(i=1,...,3^n)$  and under the assumption of linkage equilibrium between the SNPs  $p(X_i) = \prod_{i=1}^n q(x_{ij})$  with  $q(x_{ij})$  being the genotype frequency at locus *j* of the i<sup>th</sup> multilocus genotype. Under HWE for someone with 2,1 or 0 minor alleles at locus j,  $q(x_{ij}) = \alpha_j^2$ ,  $2\alpha_j(1-\alpha_j)$  or  $(1-\alpha_j)^2$ respectively, where  $\alpha_i$  is the minor allele frequency at locus j.  $S(t,X_i)$  is the survival function to age t for an individual with a multilocus genotype X, which can be expressed in terms of the baseline incidence and the logHRs according to standard survival analysis theory(2). Assuming  $S(0,X_i)=1$  for all genotypes, it is possible to solve recursively for  $\lambda_0(t)$  and hence obtain the age-specific breast cancer incidence for each multilocus genotype  $\lambda(t,X)$ . In these calculations, we used minor allele frequency estimates based on control data from population based studies (3-6) and the logHR estimates obtained from the present analysis. As an alternative, we used the estimated log(per-allele odds ratios) from the largest published population-based studies. To further evaluate the departure from the multiplicative model for the combined effects of all SNPs, we derived a risk score under the assumption of independence based on the HR estimated from the analysis of each SNP (that is, the risk score was of the form  $\exp(\beta^t X)$ ). We then estimated the HR associated with the quantiles of the distribution of the risk score and compared those against the HR predicted under the multiplicative model.

**Supplementary Figure 1 legend: A.** Cumulative distribution function of the combined HR for breast cancer risk for *BRCA2* mutation carriers based on the Odds Ratio estimates from population based studies for all nine SNPs. **B.** Predicted cumulative risk of developing breast cancer by age 80 for *BRCA2* mutation carriers by the combined HR at the above SNPs.

Supplementary Table 1: *BRCA1* and *BRCA2* mutation carriers genotyped for at least one SNP by study

Study	Country	BRCA1	BRCA2
Breast Cancer Family Registry (BCFR)	USA, Canada,	512	366
Broadt danied ranning regions (Borre)	Australia	0.2	000
Baltic Familial Breast and Ovarian Cancer Consortium	Latvia, Lithuania	102	0
(BFBOCC)	Latvia, Littidania	102	O
Copenhagen Breast Cancer Study (CBCS)	Denmark	175	88
Spanish National Cancer Centre (CNIO)	Spain, Greece	254	334
CONsorzio Studi ITaliani sui Tumori Ereditari Alla	Italy	525	325
	italy	525	323
Mammella (CONSIT TEAM)	C = 11111	60	20
Deutsches Krebsforschungszentrum (DKFZ)	Germany	69	30
HEreditary Breast and Ovarian study Netherlands	The Netherlands	796	304
(HEBON)		000	0.40
EMBRACE (TOOC)	UK and Eire	999	840
Fox Chase Cancer Centre (FCCC)	USA	82	54
German Consortium of Hereditary Breast and Ovarian	Germany	985	467
Cancer (GC-HBOC)			
Genetic Modifiers of cancer risk in BRCA1/2 mutation	France, USA	1203	599
carriers (GEMO)			
Georgetown University	USA	31	17
Gynecologic Oncology Group (GOG)	USA	390	277
Hospital Clinico San Carlos (HCSC)	Spain	132	119
Helsinki Breast Cancer Study (HEBCS)	Finland	103	105
Institut Català d'Oncologia (ICO)	Spain	112	119
International Hereditary Cancer Centre (IHCC)	Poland	350	0
Iceland Landspitali - University Hospital (ILUH)	Iceland	0	135
Interdisciplinary Health Research International	Canada	73	82
Team Breast Cancer Susceptibility (INHERIT BRCAs)	Cariada	. 0	02
Istituto Oncologico Veneto Hereditary Breast and	Italy	118	108
Ovarian Cancer Study (IOVHBOCS)	italy	110	100
kConFab	Australia	595	488
MAGIC	USA	554	313
Mayo Clinic (MAYO)	USA	243	132
Modifier Study of Quantitative Effects on Disease	Czech Republic,	324	147
(ModSQuaD)	Belgium	324	147
	USA	289	211
Memorial Sloane Kettering Cancer Center (MSKCC)		278	120
Medical University Vienna (MUV)	Austria		
National Cancer Institute (NCI)	USA	172	77
National Israeli Cancer Control Center (NICCC)	Israel	340	220
N.N. Petrov Institute of Oncology (NNPIO)	Russia	93	0
Ontario Cancer Genetics Network (OCGN)	Canada	227	172
Ohio State University Clinical Cancer Center (OSU	USA	87	49
CCG)			
Odense University Hospital (OUH)	Denmark	268	195
Pisa Breast Cancer Study (PBCS)	Italy	88	57
Sheba Medical Centre (SMC)- Tel Hashomer	Israel	614	301
Swedish Breast Cancer Study (SWE-BRCA)	Sweden	537	177
University of California Irvine (UCI)	USA	172	125
UK and Gilda Radner Familial Ovarian Cancer	USA, UK	161	35
Registries (UKGRFOCR)			
University of Pennsylvania (UPENN)	USA	313	152
Women's Cancer Research Institute (WCRI)	USA	159	69
Total		12525	7409

Supplementary Table 2: Hazard Ratio estimates after excluding prevalent cancer cases

Mutation/SNP	Unaffected/Affected	HR*	95% CI	p-value <sup>*</sup>				
SLC4A7 rs4973768								
BRCA1	4427/2528	1.04	0.97-1.13	0.25				
BRCA2	2639/1700	1.10	1.01-1.19	0.032				
STXBP4 rs6504950								
BRCA1	4466/2574	1.04	0.97-1.11	0.32				
BRCA2	2666/1720	1.02	0.92-1.12	0.74				
5p12 rs10941679								
BRCA1	4040/2430	0.99	0.92-1.07	0.82				
BRCA2	2458/1649	1.08	0.98-1.20	0.13				
BRCA2	2458/1649	1.13	1.00-1.28	0.052				
dominant								

<sup>\*</sup>Multiplicative model unless specified

Supplementary Table 3: Pairwise interaction Hazard Ratio estimates and significance tests

tests						
Locus1	Locus2	Unaffected	Breast Cancer		Interaction	
			cases			
				HR	95%CI	р-
						interaction
BRCA1						
TNRC9	2q35	3367	4117	1.09	0.95-1.24	0.22
55040						
BRCA2	TMDOO	00.40	0550	4.00	0.00.4.00	0.40
FGFR2	TNRC9	2046	2558	1.09	0.98-1.20	0.10
FGFR2	MAP3K1	2154	2706	0.94	0.85-1.03	0.17
FGFR2	LSP1	1989	2525	1.00	0.90-1.09	0.85
FGFR2	2q35	1975	2523	1.02	0.87-1.18	0.82
FGFR2	SLC4A7	1874	2408	0.98	0.89-1.07	0.64
FGFR2	5p12	1847	2406	1.04	0.92-1.19	0.53
TNRC9	MAP3K1	2127	2669	0.93	0.84-1.03	0.14
TNRC9	LSP1	1904	2408	1.04	0.66-1.63	0.86
TNRC9	2q35	1927	2465	1.11	0.94-1.30	0.22
TNRC9	SLC4A7	1823	2331	0.96	0.86-1.06	0.40
TNRC9	5p12	1806	2335	1.09	0.64-1.84	0.75
MAP3K1	LSP1	2043	2565	0.96	0.86-1.07	0.45
MAP3K1	2q35	2067	2621	1.06	0.90-1.23	0.48
MAP3K1	SLC4A7	1949	2483	1.02	0.93-1.13	0.64
MAP3K1	5p12	1925	2478	0.95	0.83-1.10	0.52
LSP1	2q35	2531	3149	1.03	0.90-1.18	0.68
LSP1	SLC4A7	2285	2888	0.92	0.84-1.01	0.07
LSP1	5p12	2213	2827	0.97	0.85-1.10	0.61
2q35	SLC4A7	2329	2961	1.10	0.96-1.26	0.17
2q35	5p12	2257	2899	0.87	0.72-1.06	0.17
SLC4A7	5p12	2562	3231	1.03	0.92-1.15	0.60
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#### Reference List

- (1) Antoniou AC, Cunningham AP, Peto J, et al. The BOADICEA model of genetic susceptibility to breast and ovarian cancers: updates and extensions. Br J Cancer 2008;98:1457-66.
- (2) Cox DR, Oakes D. Analysis of survival data. London: Chapman and Hall; 1984.
- (3) Ahmed S, Thomas G, Ghoussaini M, et al. Newly discovered breast cancer susceptibility loci on 3p24 and 17q23.2. Nat Genet 2009;41:585-90.
- (4) Easton DF, Pooley KA, Dunning AM, et al. Genome-wide association study identifies novel breast cancer susceptibility loci. Nature 2007;447:1087-93.
- (5) Milne RL, Osorio A, Cajal TR, et al. The Average Cumulative Risks of Breast and Ovarian Cancer for Carriers of Mutations in BRCA1 and BRCA2 Attending Genetic Counseling Units in Spain. Clin Cancer Res 2008;14:2861-9.
- (6) Stacey SN, Manolescu A, Sulem P, et al. Common variants on chromosome 5p12 confer susceptibility to estrogen receptor-positive breast cancer. Nat Genet 2008;40:703-6.

